What Do Market-Access Subsidies Do? Experimental Evidence from Tunisia*

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Abstract

Many countries seek to promote exports by subsidizing market access, but evidence on the effectiveness of such efforts has been mixed. We present the first randomized evaluation of a government financial-support program explicitly targeting exports, the Tasdir+ program in Tunisia. The program offered matching grants for fixed market-access costs but not variable costs. Tracking outcomes in administrative data, we find positive effects on exports on average. We find limited impacts on the number of destinations or exported products, which were stated policy targets. The finding that the fixed-cost subsidies expanded exports on the intensive margin but not the extensive margins of destinations or products stands in contrast to the predictions of several workhorse trade models.

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1 Introduction

There is growing evidence that exporting has salutary effects on firms, especially in developing countries. It can improve product quality (Verhoogen, 2008; Atkin et al., 2017; Bastos et al., 2018; Hansman et al., 2020; Demir et al., 2024), raise productivity (De Loecker, 2007; Atkin et al., 2017; Garcia-Marin and Voigtländer, 2019), induce technology adoption (Lileeva and Trefler, 2010; Bustos, 2011), increase wages (Verhoogen, 2008; Brambilla et al., 2012; Frías et al., 2024), and improve working conditions (Tanaka, 2020).

It is less clear whether government interventions to promote exports are effective. Governments dedicate substantial resources to various strategies to promote market access, including tax rebates, duty drawbacks, foreign-country trade missions, high-level delegation visits, and direct subsidies. Between 2009 and 2020, more than 2,500 export-related measures and nearly 500 initiatives providing financial assistance in foreign markets were implemented globally (Juhász et al., 2023b). But the evidence on such interventions, reviewed briefly below, is mixed. Moreover, the evaluations to date have almost uniformly relied on non-experimental methods such as matching and difference-in-difference estimators, which are subject to concerns about unobserved differences between beneficiary and non-beneficiary firms.

This paper presents the first randomized evaluation of a government financial-support program explicitly targeting exports, focusing on the Tasdir+ program in Tunisia. The program provided matching grants to offset the costs of accessing export markets. A distinctive aspect of the program is that the grants were limited to fixed costs (i.e. costs not depending on the number units produced) such as marketing expenditures, participation in trade fairs, and establishment of offices abroad. Variable costs such as salaries, materials, and transport costs were not eligible for reimbursement. The program thus offers an opportunity to study the effects of reductions specifically in fixed costs of export market access, which are commonly thought to be crucial for firms' entry into new destinations and products.

To help organize our thinking about the effects of market-access subsidies, we present a simple model of product-scope and market-penetration decisions by heterogeneous, multi-product firms, along the lines of one version of the model of Arkolakis et al. (2021), which incorporates market-penetration costs (as in Arkolakis (2010)) and

endogenous product scope (as in Eckel and Neary (2010), Bernard et al. (2011), and Mayer et al. (2014)) into a Melitz (2003)-type framework. The model generates two key implications: first, product-scope and destination-entry decisions may be insensitive to changes in fixed market-access costs over some ranges; and second, even without extensive-margin responses, sales within existing destinations and products are expected to respond to such changes.

Eligibility for the matching grants was randomized in five application rounds in 2018-2019. Pooling rounds, the randomization sample included 487 firms. The grants offered reimbursement of 50% of eligible expenditures incurred within one year as part of an approved business plan, with a cap of USD 50,000 (TND 150,000) per firm. The budget for the randomized grants was approximately USD 14 million (TND 42 million). The stated goals of the program were to increase the scale of exports and to diversify exports toward higher-value-added products and new markets.

We are able to track firms in data from several sources. We observe sales and exports from corporate tax records and employment and wages from social security records, all of which are collected in the *Repertoire National des Entreprises (RNE)* [National Repertory of Firms], a firm-level database. Importantly, this database includes exports for service firms as well as non-service firms. We also have access to transaction-level customs records, as well as administrative records from the Tasdir+ program. In addition, we conducted baseline and endline surveys.

To analyze exports, which contain many zeros, our preferred specification is a simple two-part ANCOVA with separate regressions for exporter status and for log exports among continuing exporters. We also present a Poisson Pseudo-Maximum-Likelihood (PPML) specification that combines the two margins. Using the linked customs records, we use similar specifications to examine the effects of the program on the numbers of destinations served and products exported. The study period spans the covid-19 crisis, which had a major impact in Tunisia as elsewhere. We focus primarily on outcomes in 2021, the most recent year we observe in the RNE data.

We find that the program had a positive, statistically and economically significant impact on exports. The estimates range from a 27% increase in the level of exports in the PPML specification to a 39 log point (48%) increase among continuing exporters in the ANCOVA specification. These estimates are large, but not out of line with the existing literature, as discussed below. We find positive point estimates of the effect on the

extensive margin of being an exporter, but these are not statistically significant. Using the customs records, we find little evidence that the numbers of destinations or products changed differentially between the treatment and control groups. Overall, it appears that the program achieved the goal of increasing the scale of exports, but had limited success in encouraging firms to expand the range of destinations or products exported.

Using the rich combination of datasets we have collected, we examine the mechanisms through which the program had these effects on exports. From our surveys, we find the strongest program impacts on actions to establish contracts with agents and/or distributors and to establish a foreign presence through affiliates or representatives abroad. Possibly related to this increase in foreign presence, we find that firms were more likely to become importers, although we see little increase in imports conditional on being a continuing importer. Perhaps surprisingly, we do not see significant increases in employment, earnings per employee, or wage bills of permanent employees in response to the program; we discuss possible reasons below.

We also briefly examine heterogeneity in the effects of the program by pre-program characteristics. Perhaps the most salient dimension of heterogeneity is between "totally exporting" firms — which administratively are classified in a free trade zone, outside of the Tunisian customs area — and "non-totally exporting" firms. The increase in exports we find is driven almost entirely by non-totally exporting firms. There is no evidence of an effect on exports among totally exporting firms, which were already fully integrated in international markets before the program began. There is suggestive evidence that the benefits of the program were greater for smaller and more credit-constrained firms, as well as firms planning upgrading activities such as quality certifications and product innovation, although these differences are not statistically significant.

Our findings that the subsidy for fixed market-access costs led to increases on the intensive margin of exports but not the extensive margins of destinations or products stand in contrast with the predictions of workhorse trade models in the tradition of Melitz (2003) and multi-product variants such as Bernard et al. (2011). In models of this type, fixed costs typically matter for entry, either into destinations or (when multi-product firms are considered) into products, but not for sales conditional on entry. Our results point to the importance of market-penetration costs such as those modeled by Arkolakis (2010), in which marketing expenditures in a given destination

increase the share of consumers reached.

Beyond the studies cited above, our paper is related to several strands of literature. It is most closely related to two non-experimental studies of an earlier matching-grants-for-exports program in Tunisia known as FAMEX II, Gourdon et al. (2011) and Cadot et al. (2015). Relying on matching methods, Gourdon et al. (2011) find positive effects on exports but no impacts on employment or total sales, and Cadot et al. (2015) find short-term effects on exports, the number of destinations, and the number of products, but no persistence in these effects past two years after the program. Relative to these studies, the key advantage of the current evaluation is that program eligibility was based on a lottery, rather than the judgment of program staff. Our results differ both in that we find more persistent effects on exports and that we do not find effects on the number of destinations or products.

Our paper speaks to the broader literature on government export-promotion efforts. This literature has also had to rely on non-experimental methods such as difference-in-differences and matching estimators. Among the leading studies, Volpe Martincus and Carballo (2008) find that beneficiaries of Peruvian export-promotion programs over the 2001-2005 period saw increases on the extensive margins of products and destinations but not in total exports. Álvarez and Crespi (2000) find that three export-promotion instruments in Chile (exporter committees, presence in international fairs, and utilization of business information systems) had positive effects on technological innovation but not on the number of products exported. Other notable contributions include Bernard and Jensen (2004), Görg et al. (2008), Lederman et al. (2010), Volpe Martincus and Carballo (2010a,b,c, 2012), Volpe Martincus et al. (2012), Chandra and Long (2013), Van Biesebroeck et al. (2015, 2016), Munch and Schaur (2018), Defever et al. (2020b,a), Chávez et al. (2020), Olarreaga et al. (2020), Buus et al. (2025), and Matray et al. (2024).[1] The preponderance of evidence from existing studies suggests that the effects of export-promotion programs are primarily on the extensive margin of destinations, rather than the intensive margin of exports within destinations.[2] But while the literature has generated many important insights, in the absence of

[1] Volpe Martincus (2010), Van Biesebroeck et al. (2016), and Srhoj et al. (2023) provide overviews of this literature. Relatedly, Barteska and Lee (2023) analyze the rotation of bureaucrats within the South Korean export-promotion agency and find that individual bureaucrats matter for the effectiveness of export-promotion programs.

[2] For instance, Van Biesebroeck et al. (2015, p. 1483) write, "Virtually all papers in the literature agree that the dimension of export performance most affected by these programs is the extensive margin."

randomization it is difficult to rule out the possibility that the patterns are due to unobserved differences between treated and untreated firms.

We are aware of six experimental studies related to exporting. In collaboration with an international buyer, Atkin et al. (2017) randomized initial export contracts to Egyptian rug producers and find effects on rug quality and firm productivity, but are not able to speak to the effectiveness of government interventions. Kim et al. (2018) find that informational seminars in Vietnam on average have no effect on export participation. Breinlich et al. (2017) find that informational brochures sent to firms by the UK export-promotion agency perversely made non-exporters more pessimistic about exporting. Cusolito et al. (2023) evaluate a business-training program in the Western Balkans that did not specifically target exports but that had an effect of increasing exports. The contemporaneous projects by Carvalho et al. (2024) in Brazil and Gonzalez et al. (2024) in Argentina are evaluating programs that offer information and/or consulting. Relative to these papers, the distinctive aspect of our paper is the focus on government financial support aimed explicitly at increasing exports.

There is a small experimental literature on matching grants unrelated to exporting. Several experimental evaluations have failed because of political pressures or low take-up (Campos et al., 2014). Among the few successes, McKenzie et al. (2017) find positive short-term effects of matching grants for business services in Yemen on product innovation, marketing, and adoption of accounting systems, but are not able to look at longer-term impacts because of political instability. Bruhn et al. (2018) find that a program offering subsidized consulting services to micro, small, and medium enterprises in Mexico had positive effects on sales, profits, productivity, and long-run employment. Relative to this literature, we believe that there is value in focusing on matching grants to promote exports, widely viewed as a key driver of upgrading by developing-country firms (Verhoogen, 2023).

Our paper is related to work in international trade on the role of fixed costs and the responsiveness of firms on extensive and intensive margins, outside of the context of export-promotion programs. An influential study by Das et al. (2007) estimates a dynamic model of export entry by Colombian firms and infers that the fixed (and sunk) costs of export entry are large and that subsidizing them is unlikely to be effective in increasing exports.[3] One possible reading of our results, consistent with the first

[3] Das et al. (2007) argue that this is because "exporters that need a subsidy to get into export markets

inference of Das et al. (2007), is that the matching grants were small relative to the fixed and sunk costs of entry into a new destination; that would explain the lack of response on the destination margin. But our finding that the market-access subsidies had a significant effect on the intensive margin of exports stands in contrast with the second inference. Related work on the question of whether exports expand primarily on the extensive margins of destinations and products or on the intensive margin of sales within particular destinations and products includes Kehoe and Ruhl (2013), Albornoz et al. (2023), Fitzgerald et al. (2024) and Erhardt and Gupta (2024). Relative to this literature, our study has the advantages of a clean source of exogenous variation and, specifically, variation in fixed costs as opposed to variable costs.

Finally, our paper is related to a growing literature on empirical evaluations of industrial-policy interventions, reviewed by Juhász et al. (2023a), Juhász and Steinwender (2023), Lane (2020), and Reed (2024). Relative to other forms of industrial policy, matching grants for exports have several attractive aspects: firms are required to have "skin in the game" and hence are arguably more likely to select promising investments; the subsidies are broad ("horizontal" in the terminology of Crespi et al. (2014) and others) and do not require governments to choose specific sectors or firms to support; and, as mentioned above, evidence is mounting that exporting can drive various forms of upgrading, which in turn generate positive externalities for other firms. As Reed (2024) and Juhász et al. (2023b) emphasize, export promotion has been the principal form of industrial policy in many developing countries. As one of the first randomized evaluations of an industrial-policy intervention, our paper is helping to strengthen the evidence base in this important policy area.

The paper is organized as follows. The next section presents our theoretical framework. Section 3 provides an overview of the program context and roll-out. Section 4 describes the data and estimation sample. Section 5 discusses our empirical strategy and Section 6 reports results. Section 7 concludes.

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are almost always marginal suppliers" that face relatively high entry costs, and "large incumbent exporters, who account for most of the industry's foreign sales, are unaffected by entry subsidies" (p. 868). See also Cherkashin et al. (2015) and Arkolakis et al. (2021).

2 Conceptual Framework

To guide our empirical analysis, this section presents a partial-equilibrium model of market-penetration and product-scope decisions by heterogeneous, multi-product firms. It is essentially a simplified version of the framework of Arkolakis et al. (2021) Appendix S2, which introduces market-penetration costs (à la Arkolakis (2010)) and endogenous product scope (à la Eckel and Neary (2010), Bernard et al. (2011), and Mayer et al. (2014)) into a Melitz (2003)-type framework. We provide a more fully spelled-out solution than Arkolakis et al. (2021), which we consider to be a modest theoretical contribution, but our approach is broadly consistent with theirs.[4]

2.1 Supply Side

In each source country s, there is an exogenously given continuum of potential producers of measure 1. Each has a productivity parameter, ϕ, which can be used to index firms, and a set of firm-destination-specific fixed market-access costs, c_d, as in Eaton et al. (2011), where d indexes destinations. We order a firm's products in a destination ($g = 1, 2, 3, ...G_{sd}$) by increasing distance from the firm's core competence. Following Arkolakis et al. (2021), we assume that a firm's efficiency in producing a product declines in g according to $\phi_g = \frac{\phi}{g^\alpha}$, where $\alpha \geq 0$. We normalize labor cost to 1 and allow for a destination-specific iceberg trade cost, τ_{sd}. Marginal cost is then:

$$mc_{sdg} = \frac{\tau_{sd}}{\phi_g} = \frac{\tau_{sd} g^\alpha}{\phi} \tag{1}$$

Let n_{sd} be the share of consumers in destination d that are reached by firm ϕ from source country s. As in Arkolakis et al. (2021), we assume that n_{sd} is the same for all products of a firm in a given destination. We assume that the total costs of market access are given by:

$$F_{sd}(G_{sd}, n_{sd}; c_d) = c_d \sum_{g=1}^{G_{sd}} f_{sd}(g, n_{sd}) \tag{2}$$

where for each product:

$$f_{sd}(g, n_{sd}) = \frac{g^{\delta_{sd}}}{\psi} \ln\left(\frac{1}{1 - n_{sd}}\right) \tag{3}$$

Here δ_{sd} governs how market-access costs vary with distance from core competence.

[4] Arkolakis et al. (2021) primarily focus on a version of their model without market-penetration costs.

The parameter ψ captures the cost of marketing, for instance, the cost of posting an advertisement, defined (following Arkolakis (2010)) such that a higher ψ corresponds to lower cost; we assume $\psi > 0$. Arkolakis (2010) provides a microfoundation for this specification based on a model of the visibility of advertisements.[5] The key properties are that the costs of access increase in market penetration ($\frac{\partial f_{sd}}{\partial n_{sd}} > 0$), at an increasing rate ($\frac{\partial^2 f_{sd}}{\partial n_{sd}^2} > 0$), and approach infinity as penetration approaches 1 ($\lim_{n\to 1} f_{sd} = \infty$), which guarantees less-than-full market penetration in equilibrium. Although F_{sd} varies with market penetration and product scope, it does not depend directly on firm output; in this sense, it can be thought of as representing fixed market-access costs.

2.2 Demand Side

In each destination country, there is a continuum of consumers of measure 1. Each consumer, ℓ, faces a potentially different set of firms from each source, Ω_{sd}^{ℓ}, and has constant-elasticity-of-substitution (CES) preferences over bundles of varieties offered by firms:

$$U_d = \left(\sum_{s=1}^{N} \int_{\phi \in \Omega_{sd}^{\ell}} X_{sd}(\phi)^{\frac{\sigma-1}{\sigma}} d\phi \right)^{\frac{\sigma}{\sigma-1}} \tag{4}$$

where the firm-specific bundle $X_{sd}(\phi)$ is itself a CES combination of varieties produced by the firm:

$$X_{sd}(\phi) = \left(\sum_{g=1}^{G_{sd}(\phi)} x_{sdg}(\phi)^{\frac{\sigma-1}{\sigma}} \right)^{\frac{\sigma}{\sigma-1}} \tag{5}$$

Here $x_{sdg}(\phi)$ is consumption of variety g and $G_{sd}(\phi)$ is the set of varieties offered by firm ϕ in destination d. Following Arkolakis et al. (2021), we assume that the elasticities of substitution at the variety and firm level are both equal to σ; setting the elasticities equal in the two nests simplifies the algebra but is not crucial for our predictions.[6]

Assuming a large number of firms and consumers in each country, the demand for each product in each destination is:

$$x_{sdg}(\phi) = p_{sdg}(\phi)^{-\sigma} n_{sd}(\phi) T_d {P_d}^{\sigma-1} \tag{6}$$

[5] Arkolakis (2010) allows the rate at which returns to advertising diminish, captured by his parameter β, to vary. Here we focus on the case where $\beta = 1$ in his model. This greatly simplifies the exposition without (in our view) sacrificing insights that are important for our setting.

[6] To keep the framework as simple as possible, we do not introduce an explicit term capturing product quality; note that $x_{sdg}(\phi)$ can be considered to be consumption in quality-adjusted units. See e.g. Appendix D of Kugler and Verhoogen (2012) for a discussion.

where $p_{sdg}(\phi)$ is the price of variety g offered by firm ϕ; $n_{sd}(\phi)$ is the firm's probability of reaching a given consumer; T_d is total consumer expenditure, and P_d is the aggregate CES price index corresponding to (4)-(5).[7]

2.3 Optimal Product Scope and Market Penetration

We can now consider the firm's choices of product scope and market penetration and how they respond to changes in market-access costs. To reduce clutter, we focus on the decisions of a firm from a single origin, s, selling to a single destination, d, and drop the sd subscripts and suppress the dependence on ϕ. Similar relationships hold for each source-destination combination.

Given the CES demand structure, the firm's markup over costs is constant and the optimal price is:

$$p_g = \left(\frac{\sigma}{\sigma-1}\right) mc_g = \frac{\sigma}{\sigma-1}\frac{\tau g^{\alpha}}{\phi} \tag{7}$$

Using (1), (2), (3), (6) and (7), the profit of the firm can then be written as a function of product scope, G, market penetration, n, and variables that the firm takes as exogenous:

$$\pi(G,n) = \sum_{g=1}^{G}(p_g - mc_g)x_g - c_d\sum_{g=1}^{G} f(g,n) = \Theta n\left(\sum_{g=1}^{G} g^{-\alpha(\sigma-1)}\right) - \frac{c_d}{\psi}\ln\left(\frac{1}{1-n}\right)\sum_{g=1}^{G} g^{\delta} \tag{8}$$

where $\Theta \coloneqq \left(\frac{\sigma-1}{\sigma}\frac{\phi}{\tau}\right)^{\sigma-1}\frac{T_d P_d{}^{\sigma-1}}{\sigma}$.

As in Arkolakis et al. (2021), we need to impose a restriction on the parameters to ensure that the firm's product-scope choice problem is well defined:

$$\delta + \alpha(\sigma-1) > 0 \tag{9}$$

This condition ensures that the incremental costs from adding a new product will increase more quickly than incremental revenues and hence that we will have an interior solution to the choice of product scope.[8] Below we use approximations that require additional restrictions on these parameters.

[7] That is,

$$P_d = \left(\sum_{s=1}^{N}\int_0^{+\infty}\sum_{g=1}^{G_{sd}(\phi)} p_{sdg}(\phi)^{1-\sigma} n_{sd}(\phi) h_s(\phi) d\phi\right)^{\frac{1}{1-\sigma}}$$

where $h_s(\phi)$ is the probability density of productivities in s. See Arkolakis (2010) and Arkolakis et al. (2021).

[8] This condition is an analogue in our context of the assumption in Arkolakis et al. (2021) that combined incremental scope costs are strictly increasing (their Assumption 1).

In this context, Tasdir+-type matching grants can be interpreted either as reducing destination-level market-access costs, c_d, or as lowering marketing costs within a given destination, here captured by an increase in ψ. Given our functional-form assumptions, reductions in c_d and increases in ψ are isomorphic; only the ratio c_d/ψ appears in (8). Let $a \coloneqq c_d/\psi$ represent combined market-access costs.

The fact that g and G take on only integer values makes the optimization problem non-standard — in particular, a mixed-integer-programming problem. Such problems are often difficult to solve analytically (see e.g. Boyd et al. (2007, Sec. 9.2)), and this case is no exception. To make progress analytically, we approximate the summations in (8) by Riemann integrals:[9]

$$\sum_{g=1}^{G} g^{\delta} \approx \frac{G^{1+\delta}}{1+\delta}, \qquad \sum_{g=1}^{G} g^{-\alpha(\sigma-1)} \approx \frac{G^{1-\alpha(\sigma-1)}}{1-\alpha(\sigma-1)} \tag{10}$$

which are well-defined only if the exponents in the summations are greater than -1,[10] i.e.

$$\delta > -1, \qquad \alpha(\sigma-1) < 1 \tag{11}$$

We impose these restrictions hereafter.

Using an envelope theorem with an arbitrary choice set from Milgrom and Segal (2002), optimal market penetration can be written as a function of product scope, G:

$$n^*(G) = 1 - \frac{a\Lambda}{\Theta} G^{\delta+\alpha(\sigma-1)} \tag{12}$$

where $\Lambda \coloneqq \frac{1-\alpha(\sigma-1)}{1+\delta}$ and $0 < \Lambda < 1$ by (9) and (11). See Appendix A.1.

To characterize the optimal choice of product scope, we first consider a relaxation of the problem in which product scope is treated as a continuous variable, call it $\widetilde{G}$. In Appendix

[9] These approximations are increasingly good as G grows large. To see this, let $z_g = g/G$ and note that $z_1, z_2, ..., z_G$ form a regular partition of the interval $[0,1]$. Then:

$$\lim_{G\to\infty} \sum_{g=1}^{G} g^{\delta} = G^{1+\delta} \lim_{G\to\infty} \sum_{g=1}^{G} z_g^{\delta} \triangle z_g = G^{1+\delta} \int_0^1 z^{\delta} dz = \frac{G^{1+\delta}}{1+\delta}$$

where $\triangle z_g = z_g - z_{g-1}$. A similar argument holds for $\sum_{g=1}^{G} g^{-\alpha(\sigma-1)}$.

[10] See e.g. Hunter (2014, ex. 12.25).

A.2, we derive an analytical solution for the optimal $\widetilde{G}$:

$$\widetilde{G}^* = \left[\frac{\Theta}{a\left|W_{-1}\left(-\Lambda e^{-\Lambda}\right)\right|}\right]^{\frac{1}{\delta+\alpha(\sigma-1)}} \tag{13}$$

where $W_{-1}(\cdot)$ is one branch of the Lambert W function (Corless et al., 1996).[11] Note that optimal product scope is declining in market-access costs ($\frac{\partial \widetilde{G}^*}{\partial a} < 0$). Plugging (13) into (12) gives the optimal market penetration in the relaxed case — call it $\widetilde{n}^*$ — which does not vary with market-access costs ($\frac{\partial \widetilde{n}^*}{\partial a} = 0$). (See equation (A6) in Appendix A.2.)

Turning back to the integer-constrained problem, we show in Appendix A.3 that optimal *integer* product scope, G^*, is a decreasing step function of a, defined by the following cutoffs:

$$\widehat{a}_k = \frac{\Theta H}{\Lambda|W_{-1}(-He^J)|} \tag{14}$$

where $H \coloneqq \frac{\left((k+1)^{1-\alpha(\sigma-1)}-k^{1-\alpha(\sigma-1)}\right)}{\left((k+1)^{1+\delta}-k^{1+\delta}\right)}$ and $J \coloneqq (\delta+\alpha(\sigma-1))\frac{\left((k+1)^{1+\delta}\ln(k+1)-k^{1+\delta}\ln k\right)}{(k+1)^{1+\delta}-k^{1+\delta}} - 1$. It follows from (12) that optimal market penetration in the integer-constrained problem, $n^*(a)$, is declining in a within each range $(\widehat{a}_{k+1}, \widehat{a}_k)$ and that the optimal market penetration increases discontinuously at each cutoff (i.e. $\lim_{a\to\widehat{a}_k^-} n^*(a) < \lim_{a\to\widehat{a}_k^+} n^*(a)$). It follows from (14) that the cutoff for introducing the first product in a destination is $\widehat{a}_0 = \frac{\Theta}{\Lambda}$.

The key patterns are illustrated in Figures 1 and 2. In Figure 1, the green curve represents the optimal (continuous) $\widetilde{G}^*(a)$, which declines in a; the $\{a_k\}$ are the values of a at which $\widetilde{G}^*(a)$ happens to take on integer values (i.e. at which $\widetilde{G}^*(a_k) = k$ for $k = 1, 2, 3, 4, ...$). The blue step function, with cutoffs $..., \widehat{a}_4, \widehat{a}_3, \widehat{a}_2, ...$, represents the optimal product scope under the integer constraint, $G^*(a)$.[12] In Figure 2, optimal market penetration in the relaxed problem is given by the red dotted horizontal line at $\widetilde{n}^*$. The purple sawtooth curve represents optimal market penetration under the integer constraint, $n^*(a)$, which declines in market-access costs between consecutive cutoffs $\widehat{a}_{k+1}$ and $\widehat{a}_k$.

The model thus yields two key implications. First, for a given firm, there exist ranges of a over which the extensive margins of products and destinations do not respond to a market-access subsidy (i.e. a reduction in a). In the context of Figure 1, the subsidy may not shift a across one of the cutoffs, $\widehat{a}_k$. Depending on the values of the various

[11] Given that $0 < \Lambda < 1$, $W_{-1}\left(-\Lambda e^{-\Lambda}\right) < -1$.

[12] Note that the optimal integer product scope may be above or below the optimal continuous value, depending on the value of a.

parameters, the ranges over which destination entry and destination-specific product scope remain unchanged may be wide.

Second, and perhaps more importantly, even when there are no changes in destination entry or product scope, reductions in a are expected to have an impact on the intensive margin of market penetration within a given destination-product. Between any thresholds, $\widehat{a}_{k+1}$ and $\widehat{a}_k$, a reduction in a will increase n, as illustrated by Figure 2. We see from (6) and (7) that output per destination-product, $x_g(\phi)$, will move similarly, increasing in response to the reduction in a for each $g \leq G^*$. Total output, calculated either by summing $x_g(\phi)$ for $g \in (1,2,...,G^*)$ or using the aggregator in (5), will also respond similarly. Given that markups are constant (refer to (7)), revenues will follow a similar pattern. As noted above, this prediction of an intensive-margin response of export revenues to subsidies of fixed market-access costs stands in contrast to the predictions of now-standard heterogeneous-firm models of trade such as Melitz (2003) and Bernard et al. (2011), in which fixed market-access costs matter for firms' entry into destinations and products but not for market penetration or sales conditional on entry.

Our experiment, to which we now turn, allows us to test these two key implications in a clean way.

3 Context and Experimental Design

A key motivation for Tunisian export-promotion programs has been a lack of diversification of the country's exports (Lopez-Calix et al., 2010). Approximately 75% of Tunisia's exports go to Western Europe — 58.1% to France, Italy, and Germany alone. Exports are concentrated in a small number of sectors, including machinery and transport equipment, textiles, and agricultural products.[13]

The Tunisian government has devoted substantial resources to a series of export-promotion programs, administered by its export-promotion agency, CEPEX. In 1999, it launched a matching-grants-for-exports program known as FAMEX I with a budget of USD 24 million (over multiple years). This was followed in 2006 by FAMEX II — the subject of Gourdon et al. (2011) and Cadot et al. (2015) — with a budget of USD 37.6 million. Since 1989, CEPEX has also managed a large program to defray

[13] These figures are from our own calculations using World Integrated Trade System (WITS) data.

transportation costs for Tunisian exporters, known as FOPRODEX, with a recent budget of USD 17-25 million per year.

The Tasdir+ program, known officially as the Fund for Competitiveness and Export Development Support [*Fonds d'Appui à la Compétitivité et au Développement des Exportations*], was created in 2014 with a budget of USD 23.5 million. The stated objectives were to increase exports in a sustained manner and to promote diversification towards higher-value-added exports and new markets. The program funded non-randomized waves of matching grants in 2015 and 2017.

Given the large expenditures on export promotion, the Tunisian government has been keenly interested in the cost-effectiveness of its subsidies. It decided in 2018, with encouragement from the World Bank, to randomize grants in the third and fourth waves of the Tasdir+ program.[14] The budget for the randomization sample was USD 14 million. There were four calls for applications as part of the third wave and one call as part of the fourth wave. We refer to these calls as "rounds" and focus on these five rounds hereafter.

As part of the application, firms were required to submit a "business plan" listing expenditures to be subsidized by the grant. CEPEX staff reviewed the applications to determine eligibility. To be eligible, a firm had to fulfill the following criteria: (1) be privately owned; (2) be legally based in Tunisia; (3) not be a retailer or wholesaler;[15] (4) not be an artisanal firm; (5a) (for non-agricultural firms) have a liquidity ratio, defined as assets over liabilities averaged for the three calendar years preceding the program year, greater than or equal to 1,[16] or (5b) (for agricultural firms) have five or more permanent employees or at least one export operation during the three calendar years preceding the program year; and (6) be established prior to Jan. 1, 2015 for Rounds 1-4 or Jan. 1, 2017 for Round 5. Further details on the eligibility rules are in Appendix B. Once the business plans were approved, firms were also required to respond to our baseline survey to be included in the randomization sample.

In their applications, firms could propose a budget of up TND 300,000 (USD 100,000),[17] of which the matching grant would cover 50%. Business plans could be of two types:

[14] A small number of firms received support in the third and fourth waves without being included in the randomization sample; see Appendix B for details.

[15] An exception to this rule was made for import-export firms, known locally as "trading firms."

[16] For round 5, the liquidity threshold was lowered to 0.9.

[17] The average exchange rate over our study period was approximately 3 TND/USD, and we use that rate throughout the paper.

support for exporting (*appui a l'export*) or setting up an affiliate abroad (*implantation a l'étranger*). Approximately 75% of firms chose the former. For this type of plan, eligible expenses included visits to expositions and trade fairs; market research expenditures; marketing expenditures; creation of websites and other forms of online marketing; and certification and regulatory compliance expenditures for products. For the second type of plan, eligible expenses included trademark registration in the foreign country; rent and other costs of maintaining the foreign office; travel abroad to monitor the foreign office; and technical assistance required to set up or maintain the foreign office. A full list of eligible expenditure categories appears in Table 1. Importantly for our study, variable costs such as wages and material input expenditures were not eligible for support, nor were capital investments such as purchase or installation of equipment. In their applications, firms were required to list 2-3 target export destinations, of which at least half had to be new destinations. Firms' business plans had to be approved by CEPEX staff in order for the firms to be included in the randomization sample.

Figure 3 summarizes the timeline of the experiment. The first call for applications to the experimental phase was launched in July of 2018. Randomization for the five rounds was carried out in public meetings in Sept. 2018, Nov. 2018, Feb. 2019, May 2019, and Dec. 2019. Following each randomization, the Tasdir+ steering committee formally approved the grants for selected firms and the firms were officially enrolled; this process typically took 2-3 months. Firms then had 12 months in which to incur the expenses in their business plans. Note that part of the Round 4 spending period and almost all of the Round 5 spending period coincided with the covid-19 pandemic; we return to this issue below. Control firms were excluded from re-applying to the program for at least 12 months.[18] Although we primarily rely on administrative data to track outcomes, we conducted a follow-up survey with Rounds 1-3 in July-Dec. 2020 and with Rounds 4-5 in March-Dec. 2021. When referring to pre-program information, we follow Tasdir+ practice in using 2017 as the reference year for Rounds 1-4 and 2018 for Round 5.

Within each round, randomization was stratified by sector, size, and type of business plan. Firms were classified into 6 sectors: agriculture/fishing (5% of randomization sample), trading (10%), food processing (7%), non-food manufacturing (34%), information and communication technology (ICT) services (14%), and non-ICT services (30%). Size categories for each sector (small, medium, large) were calculated based on

[18] Four control firms from Round 1 re-applied in Round 5; these firms were not included in the the Round 5 randomization sample.

the revenue distribution of previous beneficiaries of Tasdir+. In some rounds, because of prohibitively small numbers of firms in some strata, we pooled strata prior to randomization; see Appendix B for details. At the most disaggregated level, there were 28 strata.

The experiment faced a key difficulty in implementation. The original design included a "pay for performance" treatment arm in which firms that succeeded in increasing exports would receive an additional rebate, up to 40% of eligible expenses, beyond the 50% matching grant (i.e. up to 90% together). To conform to World Trade Organization (WTO) rules, these performance rebates were limited to firms selling food and agricultural goods. See Appendix B.2 for details. Approximately midway through our study, an external government auditor prohibited the Tasdir+ program from reimbursing more than 50% of expenses, effectively eliminating the rebates. In total, of the 41 firms selected to be eligible for rebates, only 5 submitted rebate requests and none received a rebate. Thus, in practice, the "Matching Grant + Rebate" arm was effectively the same as the "Matching Grant Only" arm and we treat them as a single treatment in our analysis.

Pooling rounds, 487 firms were included in the randomization sample. In Rounds 1-4, two-thirds of firms in the rebate-eligible sectors and one-half of other firms were assigned to receive matching grants. In Round 5, two-thirds of all firms were assigned to receive matching grants. Overall, among the 487 firms in the randomization sample, 281 were assigned to receive grants and 206 were assigned to control. Of the 281 firms assigned to treatment, 269 signed a contract with CEPEX and officially enrolled. For the enrolled firms, the average approved matching grant was approximately USD 30,000 (TND 90,000), based on a total business-plan budget of USD 60,000. Of the 269 enrolled firms, 187 eventually submitted at least one reimbursement request to CEPEX. Among these 187 firms, the average "realization rate," i.e. the share of the approved grant that was reimbursed by the end of the program, was 22%.[19] Thirty-five of the 187 firms that submitted at least one request received no reimbursement.

Two additional features of the institutional context are particularly relevant. First, a significant share of Tunisian firms (about 30% of our randomization sample) are

[19] A natural question is why this realization rate was so low. Conversations with firms suggest two possible explanations. One is that firms were worried that administrative snags or political inference would prevent CEPEX, the implementing agency, from actually issuing the reimbursements. A second is that there were frictions in the process of changing business plans. Although in principle firms were able to modify their plans with CEPEX approval, firms reported that in practice the process was difficult.

classified legally as "totally exporting." Administratively, these firms are in a free trade zone, outside of the Tunisian customs area. Firms in this regime sell most of their production abroad, but are allowed to sell part of their output (usually up to 30% of sales) on the domestic market.[20] Second, wages in Tunisia are in large part determined by sectoral bargaining agreements (World Bank, 2014; Angel-Urdinola et al., 2015). According to the International Labour Organization, these agreements covered around 63% of private- and public-sector workers in 2019 (ILO, 2023). The agreements stipulate base salary grids that are binding for many firms.

4 Data

This section reviews the data sources we use and the estimation samples we construct. Additional details are in Appendices C and D.

4.1 Sources

Several sources of administrative records are collected in the *Repertoire National des Entreprises (RNE)* [National Repertory of Firms], a database of registered firms collected by the Tunisian national statistical agency, *Institut National des Statistiques (INS)*. In particular, the RNE contains annual domestic sales and exports from corporate tax declarations to the *Direction Générale des Impôts (DGI)* [General Tax Authority] and firm-level employment and wages from the *Caisse Nationale de Sécurité Sociale (CNSS)* [National Social Security Fund]. We have been able to link all 487 firms from the Tasdir+ randomization sample to the RNE, although not all variables are available for all firms, as discussed below. Importantly, the RNE data contain information on exports for service firms as well as non-service firms (from the corporate tax records). We use the RNE from 2015 to 2021, the latest year to which we have access.

Through an agreement between two government ministries that we helped to facilitate, we have also been able to access data on firms' international transactions from 2017-2022 from the Tunisian customs authority. In these data, we observe transaction-level export records at the level of destination country, date of shipment,

[20] Totally exporting firms have benefited historically from an array of advantages such as VAT exemption, a reduced 10% corporate tax rate, import tax exemptions on inputs, and hiring incentives.

and 11-digit product category.[21] A shortcoming of the customs data is that firms are required to submit declarations only when they ship physical goods; service firms typically are not required to submit declarations for exported services. For this reason, we primarily rely on the RNE for information on exports, although we use the customs data when focusing on the number of destinations or products.

We also have access to administrative data from the Tasdir+ program itself, for all firms in the randomization sample. Data from firms' applications include sector, quality certifications, number of employees, sales, targeted countries, and expected effects. We also observe reimbursement requests and amounts disbursed for treated firms.

Finally, we conducted baseline and follow-up surveys. We asked about the actions firms took to increase exports and about the amounts spent on innovative activities, among other variables. Since responding to the baseline survey was a condition for eligibility for the matching grant, we have complete coverage at baseline. But response rates to the endline survey were much lower, in part due to the covid-19 pandemic and in part to the fact that CEPEX had little leverage to oblige firms to respond.[22] For this reason, we rely primarily on the information from administrative records to evaluate the impact of the program, although we present some results from the survey below.

4.2 Samples

Given the incomplete information in the RNE, customs and survey datasets, we must use different samples when analyzing different outcomes. The five samples we use are the following.

1. The *Randomization Sample* contains information from the Tasdir+ applications for all 487 firms included in the randomization.

2. The *RNE Export Sample* includes the 377 firms for which sales and exports from the DGI (tax authority) are available in the RNE in both the reference year (2017 for Rounds 1-4, 2018 for Round 5) and in 2021 (and for which sales are observed in the application data at baseline).

[21] The 11-digit categories are from the *Nomenclature de Dédouanement des Produits (NDP)* [Customs Clearance Product Nomenclature], Tunisia's most detailed product classification; the first 6 digits are harmonized with the international Harmonized System (HS) trade categories.

[22] Lower-than-usual responses during the pandemic were documented even in established surveys like the Current Population Survey in the US (Rothbaum and Bee, 2021).

3. The *RNE Employment Sample* includes the 327 firms in the *RNE Export Sample* for which the RNE also reports employment and wages from the CNSS (social security agency) in both the reference year and in 2021.

4. The *Customs Sample* contains the 210 non-service firms from the RNE Export Sample. If a non-service firm from the RNE Export Sample has no reported exports in a given year, we impute zero exports, and similarly for imports. Although service firms occasionally show up in the customs data as exporters or importers of physical goods, the customs data do not give a true picture of their engagement in international markets and we exclude them.

5. The *Survey Sample* contains the 204 firms for which we observe information on actions and spending in both the baseline and endline surveys.

4.3 Descriptive Statistics

Table 2 reports balance between the treatment and control groups for the Randomization Sample. Appendix Tables A1-A4 present similar balance tables for the other samples. To facilitate comparison, we include five key variables from the administrative data for each sample — export regime, firm age, domestic capital share, employment, and sales — in addition to variables specific to each sample. In these balance tables, as well as in the regressions below, monetary variables have been winsorized at the 3%/97% level to reduce the influence of outliers. We observe balance between the treatment and control groups, conditional on round and stratum fixed effects, in the Randomization, RNE Export, RNE Employment, and Customs Samples. At the same time, it is important to note that the patterns of missing data create differences across samples (without generating imbalance between treatment and control); in particular, firms in the RNE and Customs Samples tend to be older and larger than firms in the full Randomization Sample. Also, there is some indication of imbalance in the Survey Sample; in particular, it appears that larger (higher-sales) treated firms were more likely to respond to the endline than larger control firms. Although we do not reject the joint null of no treatment-control differences (conditional on round and stratum effects) in this sample, below we show that the results are robust to controlling for the baseline characteristics that display a lack of balance.

5 Empirical Strategy

Given that we have a randomized experiment, the empirical analysis is straightforward. The main complication is that the outcome of primary interest is exports, a variable with a significant number of zeros. There is an active econometric debate about the best way to analyze outcomes in such cases (Chen and Roth, 2024). The literature has not converged on a consensus solution. Our approach is to present estimates using what we perceive to be the two leading approaches and show that the results are robust.

Our preferred specification is a simple two-part ANCOVA specification, where we run separate regressions with (binary) exporter status and log exports as outcomes and control for values of the dependent variable in the reference year (pre-program). The basic specification is:

$$Y_{i,2021} = \beta Treated_i + \alpha Y_{i,refyear} + \gamma_r + \delta_s + \epsilon_i \tag{15}$$

where $Y_{i,2021}$ denotes outcome Y for firm i in year 2021 (e.g. either exporter status or log exports), $Y_{i,refyear}$ denotes outcome Y for firm i in the reference year (2017 for rounds 1-4, 2018 for round 5), $Treated_i$ is an indicator for being assigned to receive the matching grant in the randomization, γ_r are round fixed effects, δ_s are strata fixed effects, and ϵ_i is an error term. The approach of controlling for baseline values of the outcome has been shown to have higher power than standard difference-in-differences when autocorrelation is low (McKenzie, 2012). Note that the regression with log exports as the outcome can only be implemented among continuous exporters, and only captures the intensive margin of exports, conditional on exporting. When analyzing outcomes without zeros (for example, sales and employment), we simply estimate (15) with the log of the outcome as the dependent variable.

For outcomes with many zeros, we also present estimates from a Poisson Pseudo-Maximum-Likelihood (PPML) estimator, also know as a Poisson Quasi-Maximum-Likelihood Estimator (QMLE)) (Gourieroux et al., 1984; Santos Silva and Tenreyro, 2006; Wooldridge, 2010, section 18.2; Chen and Roth, 2024), where we include firm fixed effects to increase precision. The model is:

$$Y_{it} = \exp(\beta Treated_i \times Post_t + \gamma_i + \lambda_t)\epsilon_{it} \tag{16}$$

where Y_{it} denotes the level of the outcome (e.g. exports) for firm i in the reference year (2017 for rounds 1-4, 2018 for round 5) or 2021, $Treated_i$ is an indicator for being assigned to receive the matching grant in the randomization, $Post_t$ is an indicator for the year 2021, γ_i are firm fixed effects, λ_t are year fixed effects, and ϵ_{it} is an error term clustered by firm. In this context, $e^{\beta} - 1$ is equal to the population-average treatment effect as a proportion of the control mean. Note that this estimand combines effects on the extensive and intensive margins (e.g. of entry into exporter status and exports conditional on entry).

Our choice of outcome year merits some discussion. The RNE is an annual dataset, covering calendar years. Three of our five randomization rounds occurred in 2019 (refer to Figure 3). The first year of the covid-19 outbreak, 2020, saw a sharp contraction of economic activity in Tunisia, as elsewhere, and events in that year are difficult to interpret. We therefore focus on 2021, the latest year to which we have access in the RNE, to study post-intervention outcomes.

6 Results

6.1 Export Outcomes

Panel A of Table 3 reports two-part ANCOVA estimates of (15) in the RNE Export Sample of 377 firms. Column 1 reports a linear probability model with a 0/1 exporter indicator as the outcome. Although the point estimate is positive, indicating an increase of 5% in the probability of having positive exports, on a baseline value of 78%, we are not able to reject the null of no effect. When we focus in Column 2 on continuous exporters (i.e. the 244 firms with positive exports in both the reference year and 2021), we find a significant positive effect of 39 log points (48%).

Panel B of Table 3 reports PPML estimates of (16) for exporter status (Column 1) and the level of exports (Column 2). Although the point estimate is positive, the estimate of the treatment effect on exporter status is not significant. In Column 2, we find a marginally significant positive estimate of 0.24, equivalent to a proportional average treatment effect of (exp(0.24)-1)*100=27%. Given that this PPML estimates combines the intensive margin of exports with the extensive margin of entry into exporting, on which there is relatively little adjustment, it is not surprising that the magnitude is smaller in percentage terms that in Column 2 of Panel A.

Whichever specification we focus on, the estimated effect on exports is large in economic terms. Average exports in the reference year were approximately USD 720,000 (TND 2.15 million). The increase in exports over the 3-4 years between the reference year and 2021 is thus estimated to be on the order of USD 200,000-350,000. Recall that the program offered reimbursement of up to USD 50,000, but that the average realization rate was 22% on matching grants of approximately USD 30,000 among the 187 firms (of 281 treated firms) that submitted at least one reimbursement request. Hence the average payout was approximately US 4,400 per firm (30,000 * .22 * (187/281)). Our estimates thus suggest that the program generated an increase of USD 58-68 in exports on average for every dollar spent. This effect is large, but not out of line with the existing literature. For instance, using a non-experimental approach, it has been estimated that a USD 1 increase in the budget of trade promotion organizations is associated was an increase of approximately USD 100 in exports (ITC, 2016; Olarreaga et al., 2020; Olarreaga, 2024). In percentage terms, our estimates range from 27% to 48% over 3-4 years. Volpe Martincus and Carballo (2008) find that participation in a Peruvian export-promotion program was associated with an increase in exports of 16-36 log points (17-43%), over a shorter time period than we consider here. Volpe Martincus (2010) reports an effect size of approximately 20% for an Argentinian export-promotion program and 24% for a Colombian one.

The program appears to have been quite successful in realizing one of its primary aims, to increase Tunisian exports overall. How successful was it in realizing its other main aim, of helping firms diversify their export destinations and products? To investigate, we turn to the Customs Sample, in which we see destinations of exports and export volumes at the 11-digit product level. Table 4 reports estimates of the effect of the program on the numbers of export destinations and products. Panel A reports our ANCOVA specification, (15), and Panel B reports the PPML specification, (16). For comparison purposes, we include exporter status and exports as outcomes in Columns 1-2, which are directly comparable to Table 3, and the estimates are similar, with slightly larger standard errors (as would be expected given the reduced sample). In Columns 3-4, there is no evidence of a positive effect on the numbers of destinations or products. Indeed, the point estimates in Panel A and Column 4 of Panel B are negative.

Another way of evaluating the success of the matching grants in diversifying exports is to look at the effect on the numbers of *new* destinations or products. If the matching

grants led firms both to add new destinations and products and drop old ones, we might see no effect in total numbers of destinations and products but effects on new destinations and products. We consider a destination or product to be new if it did not appear in the customs data in the reference year for a given firm. Recall that firms had to identify 2-3 targeted destinations in their Tasdir+ applications, of which at least half had to be new destinations, so we can also examine whether firms added new targeted destinations. Table 5 presents simple regressions with different measures of new destinations or products as outcomes. With the outcomes defined in this way, there are no baseline values to control for; we present OLS estimates in Panel A and PPML estimates in Panel B, and we only include data from 2021. There is one marginally significant coefficient for the number of new targeted destinations in Column 4, Panel B. But generally the estimates reinforce the above observation that there is little evidence of an effect on the destination or product margins.

Overall, we interpret the results as suggesting that the matching grants did not meaningfully increase the diversification of treated firms' exports. Because of imprecision in the estimates, we cannot rule out modest positive effects on the order of 0.20 more destinations or 0.33 new products in response to the program. But with the standard models of Melitz (2003) and Bernard et al. (2011) in mind, one would likely have expected a larger response on these margins, especially given the large increase in exports overall. In contrast, the non-results for destinations and products accord quite naturally with the model presented above in which fixed costs matter for market penetration and entry into destinations and products may be insensitive to market-access costs over certain ranges.

6.2 Mechanisms

The rich combination of datasets we have collected allows us to explore the mechanisms underlying the reduced-form relationship between the matching grants and exports documented in Table 3. In this subsection, we consider the impacts of the matching grants on a variety of other outcomes. When the frequency of zeros is not an issue, we focus on the ANCOVA specification (equation (15)).

We first consider the effects of the program on domestic and total sales. In the presence of capacity constraints, an increase in exports induced by the program might have been expected to reduce domestic sales — the flip side of the dynamic highlighted by

Almunia et al. (2021), in which Spanish firms reacted to a domestic slump by selling more abroad. But it does not appear that this was the case. Table 6 reports both ANCOVA (equation (15)) and PPML (equation (16)) results for domestic and total sales in the RNE Export Sample. (Because some firms, particularly "totally exporting" firms, have zero domestic sales, the sample size drops in Column 2 of Panel A, when log domestic sales is the outcome.) There is little evidence of increased exports crowding out domestic sales; the point estimates for domestic sales are positive. There is some suggestive evidence of a positive effect on total sales, in particular a marginally significant estimate in the PPML results in Column 3 of Panel B, although given the noisiness of the data and the size of the standard errors, it is difficult to make definitive statements.

The grants appear not to have led firms to increase employment or wages. Table 7 considers employment and average earnings per worker at the firm level using the RNE Employment Sample (i.e. the subset of the RNE Export Sample for which we also observe employment and earnings in the reference year and 2021). For comparison purposes, Columns 1-2 report estimates for a 0/1 exporter indicator and log exports; the coefficients are similar to those in Table 3. The main new information is in Columns 3-5: we find no effects on log employment, log quarterly earnings per employee, or log total wage bills; the point estimates are -0.01, 0.01, and 0.00 respectively. Given the imprecision in our estimates, we cannot rule out modest positive effects. But we nevertheless view the lack of effects on employment and earnings to be surprising, given the substantial effects on exporting. Studies have typically found positive effects of exporting on wages (see e.g. Verhoogen (2008); Brambilla et al. (2012); Hummels et al. (2014); Frías et al. (2024)). One possible explanation relates to Tunisian labor-market institutions. As noted above, sectoral bargaining agreements are prevalent and often binding at the firm level, which may explain the lack of wage effects. Given high costs of hiring and firing in Tunisia, many firms prefer to hire workers informally or on short-term contracts (Angel-Urdinola et al., 2015). These may not show up, or may show up only partially, in the RNE employment numbers.[23] Another possible explanation is that the grants simply led to reduce slack, in line with recent research finding that many developing-country firms can increase output without significant increases in input purchases (Egger et al., 2022; Walker et al., 2024).

[23] Informal employment is not recorded in the RNE, our primary dataset, and short-term employment is recorded only if it involves a formal contract and payments to the social security agency.

Rather than increasing employment, it appears that firms' main response to the matching grants was to expand their presence abroad. In our surveys, we asked a series of questions about the actions and expenditures undertaken by firms. Table 8 reports results for several survey outcomes for the Survey Sample: an indicator for whether the firm established a new contract with a foreign distributor, agent, or partner following randomization (Column 1); an indicator for whether the firm established a new foreign affiliate or representative following randomization (Column 2); an indicator for whether the firm participated in an international fair after randomization (Column 3); and indicators for whether the firm had positive spending on certifications, new technology, travel or consulting in the previous calendar year (Columns 4-7). We simply regress the outcomes on an indicator for treatment, strata dummies and round dummies in 2021 data.[24] The statistically significant coefficients are in Columns 1-2; it appears that the main effect of the matching grants was to induce firms to strengthen their ties to destination markets, either through contracting relationships or through subsidiaries. There are suggestive positive estimates for participating in international fairs, travel and consulting, but these are not statistically significant at conventional levels. As noted above, there is some reason for concern that differential response rates to our survey between treated and control firms led to a lack of balance. To explore the robustness of the patterns, Appendix Table A6 reports specifications similar to Table 8 but where we also control for baseline sales, exporter status, and exports; the results are very similar to Table 8.

The greater presence abroad may in part be responsible for the pattern observed in Panel A of Table 9, namely that treated firms are more likely to be importers. Having a foreign partner or affiliate in a destination may also facilitate search for input suppliers from that market. However, we do not see an effect on the intensive margin of imports (Column 2 of Panel A), nor do we see an effect on the importer margin in the PPML specification (Column 1 of Panel B).

One would expect the program to have had positive effects on firms' profits. Accounting profits appear in the RNE data, and results for profits as an outcome are reported in Table 10, using the ANCOVA specification of equation (15). In Column 1, the outcome is the level of profits, in millions of 2015 dinars. In Column 2, the outcome is profits as a share

[24] The fact that the Column 1-3 outcomes are about new actions lead us to prefer this simple specification to the ANCOVA specification (equation (15)), but we also report the latter in Appendix Table A5. The results are very similar.

of total sales. Profits are notoriously difficult to measure and noisy, and the estimates are not statistically significant. But the point estimates are positive for both measures.

Overall, the results above support the interpretation that the subsidies for fixed costs of accessing foreign markets led firms to expand their marketing and customer search efforts primarily in destinations to which they were already exporting. An interesting question is how firms were able to produce more goods for export without increasing measured employment. One hypothesis is that the firms did so mainly by increasing capacity utilization and reducing slack within the firm. Another possibility is that they hired temporary workers, who do not show up in the social security records that are collected in the RNE dataset.

6.3 Heterogeneous Impacts

In this subsection, we briefly explore heterogeneity of program impacts by pre-program characteristics of firms. We focus on the main exporting outcomes using the RNE Export Sample and ANCOVA specification as in Panel A of Table 3.

Many firm characteristics tend to vary by firm size, and it is natural to ask whether the program affected smaller firms differently from larger firms. Panel A of Table 11 splits the RNE Export Sample by employment, using 50 employees as a cutoff. Effects on the extensive margin (for the 0/1 exporter variable in Columns 1 and 3) are similar. The point estimate of the effect on the intensive margin is larger for smaller firms (Column 2 vs. Column 4), although not stastically significant.

As a matching grant, the Tasdir+ program may alleviate constraints on spending especially for cash-strapped firms, allowing them to undertake investments that would otherwise be outside their budget set. To explore heterogeneity on this dimension, we calculate firms' assets-to-liabilities ratios, using information provided in their Tasdir+ applications. In Panel B of Table 11, we split the RNE Export Sample by whether firms are above or below the median of this ratio. The fact that the intensive-margin coefficient in Column 2 is larger than in Column 4 is suggestive evidence in favor of this hypothesis, although again the difference across the subsamples is not statistically significant.

As noted above, a key institutional feature in the Tunisian context is the difference between "totally exporting" firms, which are administratively outside the Tunisian

customs area, and non-totally exporting firms. Panel C of Table 11 reports separate results for these two groups. For non-totally exporting firms, we see a marginally significant response along the extensive margin of exporting and a very strong response on the intensive margin. The point estimate for the intensive margin for totally exporting firms is 0.00.[25] The program effects documented above are entirely driven by the non-totally exporting firms. It appears that the matching grants were of most help to firms that were not already fully integrated into world markets.

An advantage of having access to detailed program data is that we can explore heterogeneity by differences in the plans that firms proposed in their applications. In Panel A of Table 12, we present estimates separately by the type of business plan the firm submitted, the standard export business plan (74%) or a plan that involved setting up a foreign office. In Panel B, we consider separately firms that planned to acquire one or more quality certifications and those that had no plans for certifications. In Panel C, we present results separately for firms that had plans to introduce new products and/or tailor existing products to export destinations. There are no statistically significant differences across the split samples, but the results are suggestive that firms that implemented a standard export business plan and that had plans to upgrade by acquiring certifications or tailoring products to destinations (or introducing new ones) were better able to take advantage of the matching grants.

7 Conclusions

We have presented what we believe is the first successful randomized evaluation of a government financial-support program aimed explicitly at promoting exports. A distinctive aspect of the program is that the grants could be used only for fixed market-access costs, not variable costs. Trade theories in the tradition of Melitz (2003) and Bernard et al. (2011) predict that such subsidies will affect the extensive margins of entry into destinations and products but not the intensive margin of sales conditional on entry. Our results contrast with these predictions. We find positive effects of the program on the intensive margin of exports but little evidence of impacts on the extensive margins of destinations or products. These patterns are consistent with the theoretical model we present, a variant of the framework of Arkolakis et al. (2021), which

[25] Given that the totally exporting firms are all exporters, there is no variation in the 0/1 exporter indicator.

embeds the Arkolakis (2010) idea that fixed costs are required to reach more customers within a given destination. Because of an integer constraint on the number of products sold in a market, product scope and destination entry may be insensitive to shocks to fixed market-access costs over some ranges, but the shocks are still expected to affect market penetration for existing products and destinations.

It remains an open question how generalizable these findings are to other contexts. The non-experimental literature on export promotion has generally found stronger effects on the extensive margin of destinations than on the intensive margin of sales within destinations (Van Biesebroeck et al., 2015). What accounts for the differing results in this case? One possibility is that the random assignment is better at purging the estimates of unobservable differences between beneficiary and non-beneficiary firms than the matching and difference-in-difference estimators that have typically been employed. Another possibility is that the covid-19 pandemic, which occurred during the period of our study, made it difficult to explore new markets and encouraged firms to focus on established sales channels. A third possibility is that the specific features of the Tasdir+ program were responsible for the different impacts. In particular, to be eligible for the matching grants, firms had to submit a business plan and have it approved by Tasdir+ staff. Although in principle the plans could be changed, in practice there were frictions in doing so. Firms may have been more comfortable in developing plans for destinations and products that they were familiar with.

One hopes that further research (especially further randomized evaluations) will soon allow us to make stronger statements about which aspects of our findings are robust across contexts. But at the least, we interpret our results as providing strong evidence that fixed market-access costs *can* affect the intensive margin of market penetration, and that modeling of the behavior of firms in international markets should take this into account.

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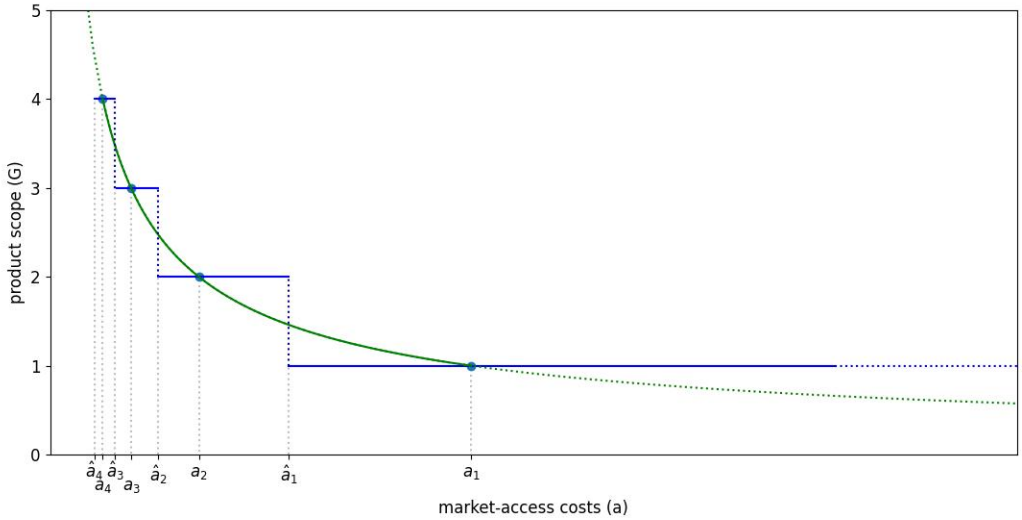
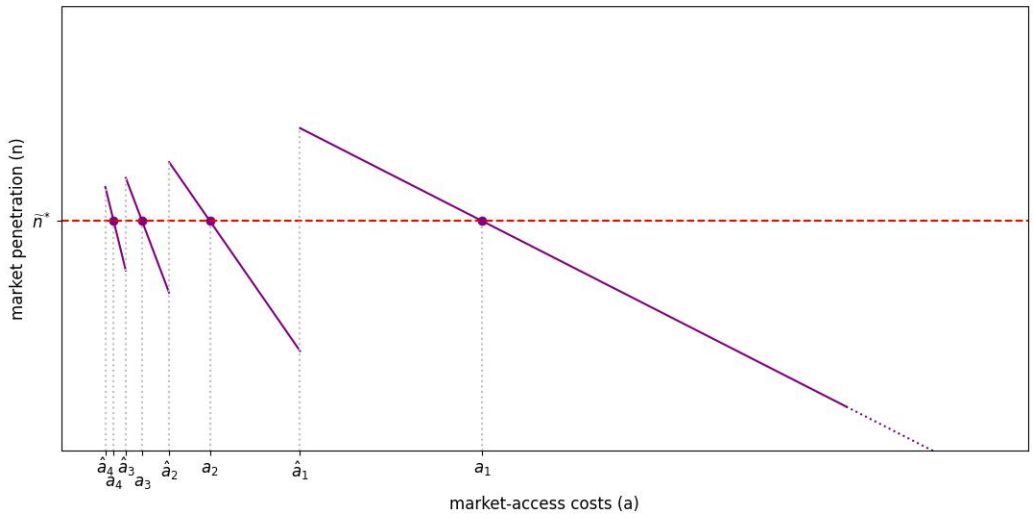
Figure 1. **Product Scope vs. Market-Access Costs**



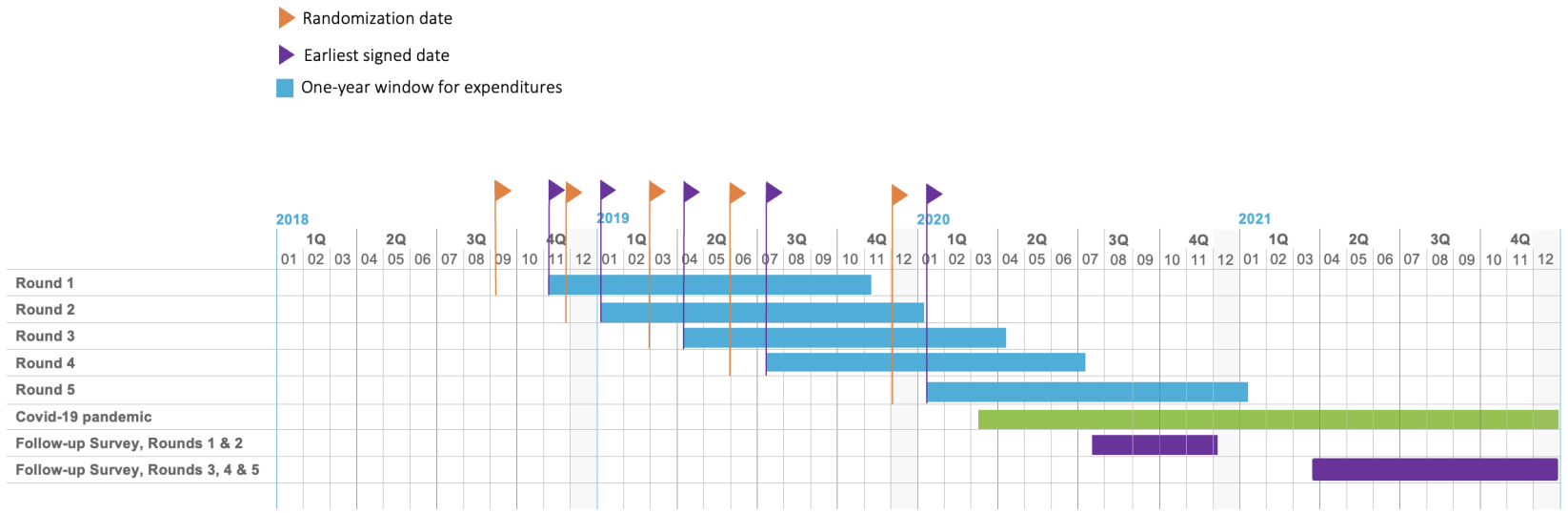
Notes: The figure is generated by a simulation of our model with the following parameter values: $\delta = 1$, $\alpha = 0.5$, $\sigma = 2$, $\Theta = 1$ (and hence $\Lambda = 0.25$). The green curve is $\widetilde{G}^*(a)$, optimal product scope in the relaxed problem (no integer constraint). The blue step function is $G^*(a)$, optimal product scope under the integer constraint. The $\{a_k\}$ are the values of a at which $\widetilde{G}^*(a)$ takes on integer values. As $a \to 0$, $k \to \infty$; for visual clarity, we omit values of $k > 4$ from the graph. The $\widetilde{G}^*(a)$ and $G^*(a)$ curves continue to the right, to the point where $a = \Theta/\Lambda$; again for visual clarity we truncate the graph.

Figure 2. **Market Penetration vs. Market-Access Costs**



Notes: The figure is generated by the same simulation as in Figure 1; see the notes to that figure. The red horizontal dotted line is the optimal market penetration in the relaxed problem (no integer constraint), $\widetilde{n}^*$, which takes on the value 0.90 in this simulation. (The y-axis extends from 0.8 to 1.0.) The sawtooth purple curve is optimal market penetration under the integer constraint, $n^*(a)$. For visual clarity, we truncate the graph as in Figure 1.

Figure 3. **Experiment Timeline**



Notes: The dates of the public randomization meetings were Sept. 6, 2018, Nov. 29, 2018, Feb. 29 2019, May 30, 2019, and Dec. 5, 2019. The baseline survey was carried out between the randomization date and the date of official enrollment. The figure reports the earliest date of enrollment for a firm from the corresponding randomization round; there was variation in the dates on which enrollment agreements were signed by firms from a given round.

Table 1. **Eligible Expenditures**

Category	Action
Distribution	Marketing/sales activities at points of sale
	Search costs in distribution channels
Firm development	Acquisition of export development software tools
	Certification fees
	Consulting in the creation of the export business plan
	Consulting in the implementation of the export business plan
	Development of skills related to exports
	Other actions related to export development
	Support for setting up an export unit
Foreign office set-up	Consulting in the creation of the export business plan
	Consulting in the implementation of the export business plan
	Incorporation costs (lawyer fees, publication, etc.)
	Missions abroad to follow-up on the office set-up
	Missions abroad to prepare the office set-up
	Promotional actions linked to establishment abroad
	Rent for office abroad
	Technical and/or commercial assistance linked to implementation
Product development	Acquisition of brands
	Adaptation of packaging to the target market(s)
	Analysis, control and testing of products intended for export
	Certification and approval of products (Halal, CE marking, etc.)
	Creation of brands
	Design/adaptation of products for target markets (prototyping, new collections)
	Registration of brands
Promotion	Creation of promotional materials
	Design and implementation of digital actions
	Organization of events, open houses etc.
	Participation in exhibits, fairs, and B2B meetings
	Website creation
	Written, audiovisual and web advertising actions
Prospecting	Broker commissions
	Development and/or acquisition of market studies
	Invitation/hosting of buyers and contractors
	Participation in calls for tenders
	Prospecting missions
	Subscription to reviews and websites
	Visits to exhibits and fairs

Notes: Translated by authors. Source is Tasdir+ internal documents, also publicized on Tasdir+ website (url: http://www.cepex.nat.tn/article/article.php?id=233).

Table 2. **Balance, Randomization Sample**

	(1) Control Mean/SD	(2) Treatment Mean/SD	(3) P-value
Totally Exporting firm	0.28	0.31	0.45
	(0.45)	(0.46)	
Age of firm (as of randomization)	14.59	14.54	0.96
	(10.78)	(10.66)	
Domestic capital share	96.89	96.66	0.90
	(13.04)	(12.70)	
Employment	39.98	44.52	0.56
	(70.52)	(114.31)	
Sales (millions 2015 dinars)	6.97	6.28	0.51
	(15.23)	(12.60)	
Exporter	0.71	0.73	0.56
	(0.46)	(0.45)	
Exports (millions 2015 dinars)	2.28	2.07	0.86
	(7.24)	(6.29)	
N	206	281	

Notes: Source is application data for reference year (2017 for Rounds 1-4, 2018 for Round 5) for Randomization Sample. (See Section 4 for details.) Standard deviations in parentheses. First row reports 0/1 indicator for being a "totally exporting" firm (administratively in a free-trade zone). Sales and exports are winsorized at the 3%/97% level. Agricultural firms were not required to report sales, export status, or exports in their applications, hence sample size is slightly smaller for these three variables (N=473, instead of 487). P-values in Column 3 are from OLS regressions of variable on treatment indicator controlling for round and stratum fixed effects. F test of joint null of no treatment-control differences (conditional on round, stratum fixed effects) for first four variables (N=487) has p-value 0.92; for all variables (N=473), the p-value is 0.92 . Monetary values were deflated to 2015 dinars using the CPI provided by INS. The average exchange rate over our study period was approximately 3 TND/USD and we use that rate throughout the paper. Additional details are in Section 4 and Appendices C and D.

Table 3. **Exports, Two-Part ANCOVA and PPML**

	Dependent variable:	
	exporter (0/1)	exports
	(1)	(2)
A. ANCOVA (exports in logs)		
Treated	0.05	0.39**
	(0.04)	(0.19)
Dep. var., refyear	0.61***	0.63***
	(0.05)	(0.05)
R2	0.46	0.65
N	377	244
Strata dummies	Y	Y
Round dummies	Y	Y
Mean of dep. var. (level)	0.66	3.92
B. PPML (exports in levels)		
Treated × Post	0.09	0.24*
	(0.06)	(0.15)
Proportional effect: $\exp(\hat{\beta}) - 1$	0.10	0.27
	(0.07)	(0.19)
Pseudo R2	0.18	0.82
N	754	754
Clusters	377	377
Firm FE	Y	Y
Year FE	Y	Y
Mean of dep. var. (level)	0.66	2.47

Notes: Sample is RNE Export Sample, for which sales and exports information (from tax records) is available in RNE in reference year and 2021. The reference year (refyear) is 2017 for Rounds 1-4 and 2018 for Round 5. Panel A reports estimates of equation (15) in text. Panel B reports Poisson Pseudo-Maximum-Likelihood (PPML) estimates of equation (16). Column 1 outcome is a 0/1 indicator for whether the firm has positive exports. Column 2 outcome is ln(exports) in Panel A, level of exports in Panel B. Means of dependent variables are for control firms at endline (2021). Column 2 dependent variable means are for the level of exports (including in Panel A). The second row of Panel B transforms the parameter estimate from first row so that it can be interpreted as a proportional effect. In Panel B, errors are clustered at the firm level. Exports are in millions of 2015 dinars. Exchange rate: 3 TND/USD. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 4. **Numbers of Destinations and Products**

	Dependent variable:			
	exporter (0/1)	exports	# destinations	# products
	(1)	(2)	(3)	(4)
A. ANCOVA (exports in logs)				
Treated	0.04	0.42*	-0.21	-0.53
	(0.05)	(0.23)	(0.34)	(1.16)
Dep. var., refyear	0.54***	0.67***	0.81***	0.78***
	(0.06)	(0.06)	(0.04)	(0.09)
R2	0.44	0.63	0.75	0.42
N	210	168	210	210
Strata dummies	Y	Y	Y	Y
Round dummies	Y	Y	Y	Y
Mean of dep. var. (level)	0.74	4.42	3.54	5.96
B. PPML (exports in levels)				
Treated × Post	0.05	0.29*	0.00	-0.09
	(0.07)	(0.17)	(0.10)	(0.21)
Proportional effect: $\exp(\hat{\beta}) - 1$	0.05	0.33	0.00	-0.09
	(0.08)	(0.22)	(0.10)	(0.19)
Pseudo R2	0.17	0.81	0.61	0.69
N	420	420	420	420
Clusters	210	210	210	210
Firm FE	Y	Y	Y	Y
Year FE	Y	Y	Y	Y
Mean of dep. var. (level)	0.74	3.64	3.54	5.96

Notes: Sample is the Customs Sample, the non-service firms in RNE Export Sample with zeros imputed for firms that do not appear in the customs export records in a given year. The reference year (refyear) is 2017 for Rounds 1-4 and 2018 for Round 5. Panel A reports estimates of equation (15) in text. Panel B reports Poisson Pseudo-Maximum-Likelihood (PPML) estimates of equation (16). Column 2 outcome is ln(exports) in Panel A and the level of exports in Panel B. # destinations is number of countries to which firm has exports. # products is number of distinct 11-digit trade categories in which firm has positive exports. Means of dependent variables are for control firms at endline (2021). Column 2 dependent variable means are for the level of exports (including in Panel A). The second row of Panel B transforms the parameter estimate from first row so that it can be interpreted as a proportional effect. In Panel B, errors are clustered at the firm level. Exports are in millions of 2015 dinars. Exchange rate: 3 TND/USD. See Section 4 and Appendices for details. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 5. **Numbers of *New* Destinations and Products**

	Dependent variable:					
	any new dest. (0/1)	# new dests.	any new targeted dest. (0/1)	# new targeted dests.	any new product (0/1)	# new products
	(1)	(2)	(3)	(4)	(5)	(6)
A. OLS						
Treated	0.03	-0.14	0.07	0.10	0.08	-0.42
	(0.07)	(0.33)	(0.06)	(0.08)	(0.07)	(1.17)
R2	0.17	0.23	0.11	0.12	0.15	0.22
N	210	210	210	210	210	210
Strata dummies	Y	Y	Y	Y	Y	Y
Round dummies	Y	Y	Y	Y	Y	Y
Mean of dep. var.	0.57	1.67	0.17	0.19	0.60	3.97
B. PPML						
Treated × Post	0.04	-0.08	0.33	0.42	0.13	-0.10
	(0.11)	(0.17)	(0.27)	(0.29)	(0.10)	(0.27)
Proportional effect: $\exp(\hat{\beta}) - 1$	0.04	-0.07	0.39	0.52	0.14	-0.10
	(0.11)	(0.16)	(0.38)	(0.44)	(0.12)	(0.25)
Pseudo R2	0.04	0.17	0.08	0.10	0.03	0.23
N	210	210	210	210	210	210
Strata dummies	Y	Y	Y	Y	Y	Y
Round dummies	Y	Y	Y	Y	Y	Y
Mean of dep. var.	0.57	1.67	0.17	0.19	0.60	3.97

Notes: Sample is the Customs Sample, the non-service firms in RNE Export Sample with zeros imputed for firms that do not appear in the customs export records in a given year. Panel A reports simple OLS regressions of dependent variable on treatment indicator, with strata and round dummies, in data for 2021. Panel B reports PPML, also with strata and round dummies in 2021 data. New destinations and products are those with positive exports in 2021 and zero exports in reference year (2017 for Rounds 1-4, 2018 for Round 5). New targeted destinations are new destinations included in list of 2-3 targeted destinations in Tasdir+ applications. A products is defined as 11-digit trade category. Means of dependent variables are for control firms at endline (2021). Column 2 dependent variable means are for the level of exports (including in Panel A). The second row of Panel B transforms the parameter estimate from first row so that it can be interpreted as a proportional effect. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 6. **Domestic & Total Sales**

	Dependent variable:		
	sells domestically	dom. sales	tot. sales
	(1)	(2)	(3)
A. ANCOVA (dom. and total sales in logs)			
Treated	0.03	0.08	0.08
	(0.02)	(0.10)	(0.07)
Dep. var., refyear	0.82***	0.73***	1.00***
	(0.02)	(0.04)	(0.04)
R2	0.85	0.86	0.90
N	377	288	377
Strata dummies	Y	Y	Y
Round dummies	Y	Y	Y
Mean of dep. var. (level)	0.79	6.70	7.66
B. PPML (dom. and total sales in levels)			
Treated × Post	0.03	0.05	0.11*
	(0.02)	(0.06)	(0.06)
Proportional effect: $\exp(\hat{\beta}) - 1$	0.03	0.05	0.12*
	(0.02)	(0.07)	(0.07)
Pseudo R2	0.19	0.88	0.85
N	754	754	754
Clusters	377	377	377
Firm FE	Y	Y	Y
Year FE	Y	Y	Y
Mean of dep. var. (level)	0.79	5.19	7.66

Notes: Sample is RNE Export Sample, for which sales and exports information (from tax records) is available in RNE in reference year and 2021. The reference year (refyear) is 2017 for Rounds 1-4 and 2018 for Round 5. Panel A reports estimates of equation (15) in text. Panel B reports Poisson Pseudo-Maximum-Likelihood (PPML) estimates of equation (16). Column 1 outcome is a 0/1 indicator for whether the firm has positive domestic sales. Column 2-3 outcomes are log domestic and total sales in Panel A, levels of domestic and total sales in Panel B. Means of dependent variables are for control firms at endline (2021). Columns 2-3 dependent variable means are for the level of domestic and total sales (including in Panel A). The second row of Panel B transforms the parameter estimate from first row so that it can be interpreted as a proportional effect. In Panel B, errors are clustered at the firm level. Domestic and total sales are in millions of 2015 dinars. Exchange rate: 3 TND/USD. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 7. **Employment**

	Dependent variable:				
	exporter (0/1)	ln(exports)	ln(emp.)	ln(avg qtr earnings)	ln(wage bill)
	(1)	(2)	(3)	(4)	(5)
Treated	0.04	0.40*	-0.01	0.01	0.00
	(0.04)	(0.20)	(0.06)	(0.03)	(0.07)
Dep. var., refyear	0.61***	0.67***	0.87***	0.76***	0.85***
	(0.05)	(0.05)	(0.03)	(0.04)	(0.03)
R2	0.48	0.66	0.90	0.75	0.88
N	327	209	327	327	327
Strata dummies	Y	Y	Y	Y	Y
Round dummies	Y	Y	Y	Y	Y
Mean of dep. var. (level)	0.66	4.18	64.60	11.14	712.73

Notes: Sample is RNE Employment Sample, the subset of RNE Export Sample for which employment and earnings (from social security agency) are available in RNE in reference year and 2021. The reference year (refyear) is 2017 for Rounds 1-4 and 2018 for Round 5. Table reports ANCOVA estimates of equation (15) in text. Means of dependent variables are for control firms at endline (2021). Columns 2-5 dependent variable means are for levels (not logs). Earnings and wage bill are in thousands of 2015 dinars. Exchange rate: 3 TND/USD. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 8. **Survey Outcomes**

	Dependent variable:						
	new contract with foreign dist./ agent/partner	new foreign affiliate/ representative	participated in int'l fair	spent on certifications	spent on new tech.	spent on travel	spent on consulting
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Treated	0.12*	0.09**	0.07	0.03	0.01	0.07	0.07
	(0.06)	(0.04)	(0.06)	(0.06)	(0.07)	(0.07)	(0.07)
R2	0.32	0.34	0.29	0.22	0.22	0.24	0.23
N	204	204	204	204	204	204	204
Strata dummies	Y	Y	Y	Y	Y	Y	Y
Round dummies	Y	Y	Y	Y	Y	Y	Y
Mean of dep. var.	0.24	0.06	0.23	0.19	0.54	0.57	0.30

Notes: Sample is Survey Sample, omitting round 5 firms. Table reports OLS estimates of simple regressions of dependent variable on treatment indicator, strata and round dummies. Dependent variable in Column 1 is indicator for whether firm contracted with distributor, local agent or partner in foreign market since randomization. In Column 2, it is indicator for having established a new foreign affiliate or representative. In Column 3, it is indicator for having participated in international fair/expo. In Columns 4-7, they are indicators for having positive spending in indicated category in calendar year 2019. ANCOVA specifications including the dependent variable at baseline are reported in Appendix Table A5. Means of dependent variables are for control firms at endline. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 9. **Imports**

	Dependent variable:	
	importer (0/1)	imports
	(1)	(2)
A. ANCOVA (imports in logs)		
Treated	0.09**	-0.02
	(0.04)	(0.18)
Dep. var., refyear	0.50***	0.85***
	(0.06)	(0.05)
R2	0.61	0.78
N	210	161
Strata dummies	Y	Y
Round dummies	Y	Y
Mean of dep. var. (level)	0.79	9.30
B. PPML (imports in levels)		
Treated × Post	0.07	-0.07
	(0.05)	(0.12)
Proportional effect: $\exp(\hat{\beta}) - 1$	0.08	-0.07
	(0.05)	(0.11)
Pseudo R2	0.11	0.86
N	494	494
Clusters	255	255
Firm FE	Y	Y
Year FE	Y	Y
Mean of dep. var. (level)	0.81	6.38

Notes: Sample is the Customs Sample, the non-service firms in RNE Export Sample with zeros imputed for firms that do not appear in the customs export records in a given year. The reference year (refyear) is 2017 for Rounds 1-4 and 2018 for Round 5. Panel A reports estimates of equation (15) in text. Panel B reports Poisson Pseudo-Maximum-Likelihood (PPML) estimates of equation (16). Column 2 outcome is ln(imports) in Panel A and the level of imports in Panel B. Means of dependent variables are for control firms at endline (2021). Column 2 dependent variable means are for the level of imports (including in Panel A). The second row of Panel B transforms the parameter estimate from first row so that it can be interpreted as a proportional effect. In Panel B, errors are clustered at the firm level. Imports are in millions of 2015 dinars. Exchange rate: 3 TND/USD. See Section 4 and Appendices for details. Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 10. **Profits**

	Dependent variable:	
	profit (level)	profit/sales
	(1)	(2)
Treated	0.19	0.03
	(0.12)	(0.03)
Dep. var., refyear	0.43***	0.45***
	(0.05)	(0.06)
R2	0.31	0.23
N	341	341
Strata dummies	Y	Y
Round dummies	Y	Y
Mean of dep. var.	0.40	0.02

Notes: Sample is firms in RNE Export Sample for which profits are reported both in the reference year and 2021. Table reports ANCOVA estimates of equation (15) in text. Column 1 dependent variable is level of reported profits, in millions of 2015 dinars. Column 2 dependent variable is ratio of profits to sales in same year. The reference year is 2017 for Rounds 1-4 and 2018 for Round 5. Means of dependent variables are for control firms at endline (2021). Additional details are in Section 4 and Appendices C and D. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 11. **Heterogeneity by Size, Financial Condition, Trade Regime**

	exporter (0/1) (1)	ln(exports) (2)	exporter (0/1) (3)	ln(exports) (4)
A. By Initial Employment				
	<50 employees		>50 employees	
Treated	0.05	0.48*	0.04	0.34
	(0.05)	(0.26)	(0.06)	(0.27)
Dep. var., refyear	0.63***	0.70***	0.54***	0.53***
	(0.06)	(0.07)	(0.09)	(0.06)
R2	0.47	0.67	0.52	0.53
N	240	139	137	105
Mean of dep. var. (level)	0.58	0.96	0.80	7.59
B. By Financial Condition				
	Low assets/liabilities		High assets/liabilities	
Treated	0.05	0.66**	0.07	0.33
	(0.06)	(0.26)	(0.06)	(0.31)
Dep. var., refyear	0.52***	0.69***	0.65***	0.61***
	(0.07)	(0.08)	(0.07)	(0.07)
R2	0.44	0.72	0.54	0.68
N	201	125	176	119
Mean of dep. var. (level)	0.66	3.46	0.65	4.42
C. By Trade Regime				
	Non-totally exporting		Totally exporting	
Treated	0.08*	0.71**		0.00
	(0.05)	(0.28)		(0.20)
Dep. var., refyear	0.52***	0.55***		1.00***
	(0.06)	(0.07)		(0.09)
R2	0.43	0.61		0.87
N	277	144		100
Mean of dep. var. (level)	0.54	2.72		5.66

Notes: Table reports regressions similar to Panel A of Table 3 using the RNE Export Sample split along the indicated dimensions. In Panel B, sample is split into with below or above-median values of assets-to-liabilities ratios in reference year. All regressions include strata and round dummies. Means of dependent variables are for control firms at endline (2021). Columns 2 and 4 dependent variable means are for the level of exports. Exports are in millions of 2015 dinars. Exchange rate: 3 TND/USD. Additional details are in Section 4 and Appendices C and D. T tests of null of no coefficient differences for exporter have p-values 0.98 and 0.52 for Panels A and B, respectively; for ln(exports), they have p-values 0.61, 0.35, and 0.01 for Panels A, B, and C, respectively. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 12. **Heterogeneity by Business Plan Characteristics**

	exporter (0/1) (1)	ln(exports) (2)	exporter (0/1) (3)	ln(exports) (4)
A. By Type of Business Plan				
	Affiliate abroad		Support for exporting	
Treated	0.02	0.18	0.05	0.46**
	(0.07)	(0.49)	(0.04)	(0.20)
Dep. var., refyear	0.72***	0.44***	0.56***	0.75***
	(0.08)	(0.10)	(0.06)	(0.06)
R2	0.68	0.70	0.42	0.68
N	82	55	295	189
Mean of dep. var. (level)	0.67	4.44	0.66	3.76
B. By Plan to Spend on Certifications				
	No certification plan		Has certification plan	
Treated	0.07	0.24	0.06	0.40
	(0.05)	(0.24)	(0.08)	(0.37)
Dep. var., refyear	0.59***	0.63***	0.63***	0.62***
	(0.06)	(0.06)	(0.10)	(0.10)
R2	0.49	0.66	0.48	0.72
N	265	170	112	74
Mean of dep. var. (level)	0.65	3.48	0.68	5.10
C. By Plan to Spend on Product Tailoring/Innovation				
	No tailoring/innovation plan		Has tailoring/innovation plan	
Treated	0.07	0.22	0.02	0.54
	(0.05)	(0.22)	(0.07)	(0.40)
Dep. var., refyear	0.60***	0.62***	0.64***	0.68***
	(0.06)	(0.06)	(0.08)	(0.10)
R2	0.48	0.73	0.55	0.56
N	259	163	118	81
Mean of dep. var. (level)	0.62	4.21	0.76	3.30

Notes: Table reports regressions similar to Panel A of Table 3 using the RNE Export Sample split along the indicated dimensions. In Panel A, sample is split by type of business plan submitted to Tasdir+ as part of application. In Panel B, sample is split by whether firm planned expenditures on certifications as part of business plan. In Panel B, sample is split by whether firm planned expenditures on adapting products to destination market (or creating new products). Means of dependent variables are for control firms at endline (2021). Columns 2 and 4 dependent variable means are for the level of exports. Exports are in millions of 2015 dinars. Exports are in millions of 2015 dinars. Exchange rate: 3 TND/USD. Additional details are in Section 4 and Appendices C and D. T tests of null of no coefficient differences for exporter have p-values 0.95, 0.78, and 0.73 for Panels A, B, and C, respectively; for ln(exports), they have p-values 0.46, 0.61, and 0.35 for Panels A, B, and C, respectively. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.